Optimization of ARTM regularization coefficients via stochastic variational inference

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Outline

- 1. Topic modeling
 - 1.1 ARTM probabilistic model
- 2. Evidence maximization
 - 2.1 Variational inference
 - 2.2 Reparametrization trick for posterior
 - 2.3 Logistic-normal distribution
 - 2.4 How to optimize with normalizing constant of prior unknown?
- 3. Review

Topic modeling

Topic modeling an algorithmic tool that help us organize, search and understand vast amount of information.

Applications

- ► Exploratory search
- Clustering
- Creating annotations
- ► Recomendation systems
- ▶ Looking for similar documents in a huge collection
- .

Probabilistic Topic Modeling

Topic proportions and **Topics Documents** assianments gene 0.04 0.02 dna Seeking Life's Bare (Genetic) Necessities genetic 0.01 COLD SPRING HARBOR, NEW YORK-"are not all that far apart," especially in How many genes does an organism need to comparison to the 75,000 genes in the husurvive! Last week at the genome meeting many genome, notes Siv Andersson or Signature. here, two genome researchers with radically University in different approaches presented complementary views of the basic genes needed for life life 0 02 evolve 0.01 One research team, using computer analyses to compare known genomes, concluded organism 0.01 that today's organisms can be sustained with sequenced. "It may be a way of organ just 250 genes, and that the earliest life forms any newly sequenced genome," explains required a mere 128 genes. The Arcady Mushegian, a computational moother researcher marved genes lecular biologist at the National Center in a simple parasite and estifor Biotechnology Information mated that for this organism. in Bethesda, Maryland, Comparine 800 genes are plenty to do the 0.04 brain iob-but that anything short 0.02 neuron of 100 wouldn't be enough. nerve 0.01 Although the numbers don't match precisely, those predictions * Genome Mapping and Sequencing, Cold Spring Harbor, New York, Stripping down. Computer analysis yields an esti-May 8 to 12. mate of the minimum modern and ancient genomes. data 0.02 SCIENCE • VOL. 272 • 24 MAY 1996 number 0.02 computer 0.01

Probabilistic Topic Modeling

Consider a set of documents D, dictionary W. Each document d consists of some words $w \in W$.

- ▶ There's a set of latent topics T. Each $t \in T$ is a distribution over W: $\varphi_{wt} = p(w|t)$
- ▶ Each document $d \in D$ has specific distribution over T: $\theta_{td} = p(t|d)$.
- Bag of words assumption: order of words in d doesn't matter. Words in d are i.i.d. from p(w|d)
- ▶ Conditional independence assumption: p(w|d, t) = p(w|t).
- ▶ Law of total probability: $p(w|d) = \sum_{t} \underbrace{p(w|t)}_{\varphi_{wt}} \underbrace{p(t|d)}_{\theta_{td}}$

- ▶ $N = (n_{dw})_{W \times D}$ matrix of counts of words in documents
- ▶ $F = (n_{dw}/n_d)_{W \times D}$ normalized matrix of counts; $n_d \equiv \sum_w n_{dw}$.
- $(\varphi_{wt})_{W \times T} \equiv \Phi \in \mathbb{R}^{W \times T}$ topic-specific distribution over words.
- $(\theta_{td})_{T \times D} \equiv \Theta \in \mathbb{R}^{T \times D}$ document-specific distribution over topics.

$$p(w|d) = \sum_{t} \varphi_{wt} \theta_{td}$$

Maximization of log-likelihood:

$$\mathit{p}(\mathit{N}|\Phi,\Theta) = \prod_{d} \prod_{w} (\sum_{t} \varphi_{wt} \theta_{td})^{n_{dw}} \rightarrow \max_{\Phi,\Theta}$$

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$$\begin{split} p(\textit{N}|\Phi,\Theta) &= \prod_{d} \prod_{w} (\sum_{t} \varphi_{wt} \theta_{td})^{n_{dw}} \rightarrow \max_{\Phi,\Theta} \\ \log p(\textit{N}|\Phi,\Theta) &= \sum_{d} \sum_{w} n_{dw} \log \sum_{t} \varphi_{wt} \theta_{td} \rightarrow \max_{\Phi,\Theta} \end{split}$$

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Non-negative matrix factorization: $F \approx \Phi \Theta$:

$$\begin{split} \mathcal{L}(\Phi,\Theta) &\equiv \log p(\textit{N}|\Phi,\Theta) \rightarrow \max_{\Phi,\Theta} \\ \mathcal{L}(\Phi,\Theta) &= -\sum_{d} n_{d} \operatorname{KL}\left(\textit{F}_{d} \mid\mid (\Phi\Theta)_{d}\right) \end{split}$$

III-posed problem:
$$F = \Phi\Theta = (\Phi S)(S^{-1})\Theta)$$

Prior on Φ, Θ :
$$\hat{p}(\Phi, \Theta | \tau) \propto \exp(\sum_i \tau_i R_i(\Phi, \Theta))$$

$$p(N, \Phi, \Theta | \tau) = p(N | \Phi, \Theta) \hat{p}(\Phi, \Theta | \tau) \rightarrow \max_{\Phi, \Theta}$$

Regularized non-negative matrix factorization: $F \approx \Phi\Theta$:

$$\begin{split} \underbrace{\mathcal{L}(\Phi,\Theta)}_{\text{log-likelihood}} + \underbrace{\textit{\textbf{R}}(\Phi,\Theta)}_{\text{regularization}} &\rightarrow \max_{\Phi,\Theta} \\ \mathcal{L}(\Phi,\Theta) = -\sum_{d} n_{d} \operatorname{KL}\left(F_{d} \mid\mid (\Phi\Theta)_{d}\right) \\ \textit{\textbf{R}}(\Phi,\Theta) = \sum_{i} \tau_{i} R_{i}(\Phi,\Theta) \end{split}$$

 $R_i(\Phi,\Theta)$ — any differentiable regularization function

ARTM — regularizers

We can use $R(\Phi, \Theta)$ to:

- Make model better (easy to interpret, coherent topics)
 - Smoothing/sparsing (LDA)

$$R = \sum_{t \in T} \sum_{w \in W} \beta_{wt} \log \varphi_{wt} + \sum_{d \in D} \sum_{t \in T} \alpha_{td} \log \theta_{td}$$

- ▶ Decorellation: $R = -\tau \sum_{t \in T} \sum_{s \in T \setminus t} \sum_{w \in W} \varphi_{ws} \varphi_{ws}$
- Add soft restrictions to the model
 - Semi-supervised regularizer

$$\textit{R} = \sum_{\textit{t} \in \mathcal{T}_0} \sum_{\textit{w} \in \textit{W}_\textit{t}} \beta_{\textit{wt}} \log \varphi_{\textit{wt}} + \sum_{\textit{d} \in \textit{D}_0} \sum_{\textit{t} \in \mathcal{T}_\textit{d}} \alpha_{\textit{td}} \log \theta_{\textit{td}}$$

- Add new information to the model
 - Author-topic models
- Make the model specified for solving specific problem
 - Topic models with time
 - Topic models for classification
- **.**..

ARTM — advantages

- ▶ Really easy to infer new models
- ▶ There is an iterative algorithm for any differentiable $R(\Phi, \Theta)$ based on fixed-point iteration method for stationarity conditions
- You can combine any amount of regularizers in one model

ARTM — problems

How to fit hyperparamters (regularization coefficients τ)?

Current approach: heuristic, manual fitting. How to fit them automatically is an open problem

- ► Too many parameters for grid-search
- We don't know what to optimize

Maximum-evidence approach

$$p(N| au) = \int p(N|\Phi,\Theta)p(\Phi,\Theta| au)d\Phi d\Theta \longrightarrow \max_{ au}$$

Or, if we denote (Φ, Θ) as \mathbf{z} , and \mathbf{N} as \mathbf{x}

$$p(\mathbf{x}|\mathbf{\tau}) = \int p(\mathbf{x}|\mathbf{z})p(\mathbf{z}|\mathbf{\tau})d\mathbf{z} = \mathbb{E}_{\mathbf{z}|\mathbf{\tau}}p(\mathbf{x}|\mathbf{z}) \longrightarrow \max_{\mathbf{\tau}}$$

We'll use variational infference to optimize it.

Variational inference

Recall our probabilistic model:

$$\begin{split} p(\mathbf{x}|\mathbf{z}) &\equiv p(N|\Phi,\Theta) = \prod_{d} \prod_{w} (\sum_{t} \varphi_{wt} \theta_{td})^{n_{dw}} \\ p(\mathbf{z}|\mathbf{\tau}) &\equiv p(\Phi,\Theta|\mathbf{\tau}) \propto \exp(\sum_{i} \tau_{i} R_{i}(\Phi,\Theta)) \end{split}$$

We introduce some family of distributions $q(\mathbf{z}|\mathbf{\lambda}) \approx p(\mathbf{z}|\mathbf{x})$

$$\begin{split} &\log p(\mathbf{x}|\boldsymbol{\tau}) \geq \mathcal{F}(\boldsymbol{\lambda}, \boldsymbol{\tau}) \equiv \int q(\mathbf{z}|\boldsymbol{\lambda}) \log \frac{p(\mathbf{x}, \mathbf{z}|\boldsymbol{\tau})}{q(\mathbf{z}|\boldsymbol{\lambda})} d\mathbf{z} = \\ &= \mathbb{E}_{q_{\lambda}} \log p(\mathbf{x}|\mathbf{z}) - \mathbb{E}_{q_{\lambda}} \log \frac{q(\mathbf{z}|\boldsymbol{\lambda})}{\hat{p}(\mathbf{z}|\boldsymbol{\tau})} - \log Z(\boldsymbol{\tau}) \rightarrow \max_{\boldsymbol{\tau}, \boldsymbol{\lambda}} \end{split}$$

 $\mathcal{F}(\lambda, au)$ — ELBO (evidence lower-bound) Variational inference can be used to optimize evidence over hyperparamters of the model.

Reparametrization trick

Stochastic optimization problem:

$$egin{aligned} oldsymbol{z} &\sim p(oldsymbol{z}) \ \mathbb{E}_{oldsymbol{z}} f(oldsymbol{z}, oldsymbol{\lambda}) &
ightarrow \max_{oldsymbol{\lambda}} \
abla_{oldsymbol{\lambda}} \mathbb{E}_{oldsymbol{z}} f(oldsymbol{z}, oldsymbol{\lambda}) = ? \end{aligned}$$

$$abla_{m{\lambda}} \mathbb{E}_{m{z}} \mathit{f}(m{z}, m{\lambda}) = \mathbb{E}_{m{z}}
abla_{m{\lambda}} \mathit{f}(m{z}, m{\lambda}) pprox rac{1}{m} \sum_{i=1}^{m}
abla_{m{\lambda}} \mathit{f}(m{z}_{m{i}}, m{\lambda})$$

We can use stochastic gradient descent

Reparametrization trick

What if probability p(z) depends on λ ?

$$\mathbf{z} \sim p(\mathbf{z}|\mathbf{\lambda}) \equiv p_{\lambda}(\mathbf{z})$$

$$\mathbb{E}_{p_{\lambda}} f(\mathbf{z}) \to \max_{\mathbf{\lambda}}$$

$$\nabla_{\mathbf{\lambda}} \mathbb{E}_{\mathbf{z}} f(\mathbf{z}) = ?$$

Reparametrization trick

Let's assume that if $\mathbf{z} \sim q(\mathbf{z}|\boldsymbol{\lambda}) \equiv q_{\lambda}$, $\mathbf{z} = g(\boldsymbol{\varepsilon}, \boldsymbol{\lambda})$, where $\varepsilon \sim q_0$

$$\nabla_{\lambda} \mathbb{E}_{q_{\lambda}} f(\mathbf{z}) = \mathbb{E}_{q_{0}} \nabla_{\lambda} f(g(\boldsymbol{\varepsilon}, \boldsymbol{\lambda})) \approx \frac{1}{m} \sum_{i=1}^{m} \nabla_{\lambda} f(g(\boldsymbol{\varepsilon}_{i}, \boldsymbol{\lambda}))$$

Example:

$$\begin{split} \mathbf{z} &\sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}); \quad \boldsymbol{\lambda} \equiv (\boldsymbol{\mu}, \boldsymbol{\Sigma}) \\ g(\boldsymbol{\varepsilon}, \boldsymbol{\lambda}) &= \boldsymbol{\mu} + \boldsymbol{\Sigma}^{1/2} \boldsymbol{\varepsilon}; \quad \boldsymbol{\varepsilon} \sim \mathcal{N}(0, \mathbf{I}) \\ \nabla_{\boldsymbol{\lambda}} \mathbb{E}_{\boldsymbol{q}_{\boldsymbol{\lambda}}} \mathit{f}(\mathbf{z}) &= \mathbb{E}_{\boldsymbol{q}_{\boldsymbol{0}}} \mathit{f}(\boldsymbol{\mu} + \boldsymbol{\Sigma}^{1/2} \boldsymbol{\varepsilon}) \end{split}$$

Estimation has low variance.

Gradients of ELBO

Reparametrization trick for $q(\mathbf{z}|\lambda)$:

$$\begin{aligned} \mathbf{z} &= g(\varepsilon, \boldsymbol{\lambda}); \quad \varepsilon \sim q_0 \\ \mathcal{F}(\boldsymbol{\lambda}, \boldsymbol{\tau}) &= \mathbb{E}_{q_{\lambda}} \log p(\mathbf{x}|\mathbf{z}) - \mathbb{E}_{q_{\lambda}} \log \frac{q(\mathbf{z}|\boldsymbol{\lambda})}{\hat{p}(\mathbf{z}|\boldsymbol{\tau})} - \log Z(\boldsymbol{\tau}) \\ \frac{\partial \mathcal{F}}{\partial \boldsymbol{\lambda}} &= \mathbb{E}_{q_0} \nabla_{\boldsymbol{\lambda}} \log p(\mathbf{x}|g(\varepsilon, \boldsymbol{\lambda})) - \mathbb{E}_{q_0} \nabla_{\boldsymbol{\lambda}} \log q(g(\varepsilon, \boldsymbol{\lambda})|\boldsymbol{\lambda}) \\ &+ \mathbb{E}_{q_0} \nabla_{\boldsymbol{\lambda}} \log \hat{p}(g(\varepsilon, \boldsymbol{\lambda})|\boldsymbol{\tau}) \\ \frac{\partial \mathcal{F}}{\partial \boldsymbol{\tau}} &= \mathbb{E}_{q_0} \nabla_{\boldsymbol{\tau}} \log \hat{p}(g(\varepsilon, \boldsymbol{\lambda})|\boldsymbol{\tau}) - \mathbb{E}_{q_0} \nabla_{\boldsymbol{\tau}} \log Z(\boldsymbol{\tau}) \end{aligned} \tag{1}$$

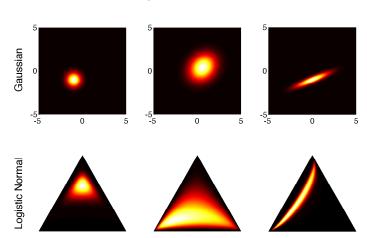
Algorithm

Algorithm 1 Sketch

```
Data: Flow of data \mathbf{x}; learning rate \kappa Result: \lambda, \tau initialize \lambda and \tau repeat \mathbf{x} \leftarrow \{\text{Get mini-batch}\} Sample \varepsilon \leftarrow q_0 Compute stochastic estimations of \frac{\partial \mathcal{F}}{\partial \tau} and \frac{\partial \mathcal{F}}{\partial \lambda} according to (1) \tau = \tau - \kappa \frac{\partial \mathcal{F}}{\partial \tau} // or any other stochastic-gradient-kind \lambda = \lambda - \kappa \frac{\partial \mathcal{F}}{\partial \lambda} // optimization algorithm, e.g. AdaGrad until convergence
```

Logistic-normal distribution

$$\begin{split} \boldsymbol{\varepsilon} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}); \quad \boldsymbol{z} &= \mathcal{P}(\boldsymbol{\varepsilon}) \implies \boldsymbol{z} \sim \mathcal{P}(\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})), \\ \text{where } \mathcal{P}(\boldsymbol{x}) &= \boldsymbol{z} \implies \boldsymbol{z}_i = \frac{\exp(\boldsymbol{x}_i)}{\sum_i \exp(\boldsymbol{x}_i)} \text{— softmax function} \end{split}$$



Logistic-normal distribution — Properties

- ▶ Distribution over simplex $S_x = \{x : x_i \ge 0; \sum_i x_i = 1\};$
- Similar to Dirichilet, but never exactly the same;
- \blacktriangleright Although can be approximated with Dirichilet with large α
- ▶ No analytical solution for mode/mean/variance.
- $lackbox{ We can write down component-wise median: } \mathcal{P}(oldsymbol{\mu})$
- Easy to sample from

Reparametrization of q

$$\begin{split} \mathbf{z} &= \mathbf{g}(\boldsymbol{\varepsilon}, \boldsymbol{\lambda}) = \mathcal{P}(\mathrm{diag}(\boldsymbol{\sigma})\boldsymbol{\varepsilon} + \boldsymbol{\mu}); \quad \boldsymbol{\varepsilon} \sim \mathcal{N}(0, \mathrm{I}) \\ \mathbf{z} \sim \mathcal{P}(\mathcal{N}(\boldsymbol{\mu}, \mathrm{diag}(\boldsymbol{\sigma}))) \end{split}$$

where λ denotes (μ, σ) .

Individual vector of parameters $\pmb{\mu}$ and $\pmb{\sigma}$ for each column of Φ and Θ

$$\log q(g(\varepsilon, \lambda)) = \log q_0(\varepsilon) + \log \left| \frac{\partial g}{\partial \varepsilon} \right| = \log q_0(\varepsilon) + \sum_k (\log z_k + \log \sigma_k)$$

Recall gradients of ELBO

$$\begin{split} & \mathbf{z} = \mathbf{g}(\boldsymbol{\varepsilon}, \boldsymbol{\lambda}); \quad \boldsymbol{\varepsilon} \sim \mathcal{N}(0, \mathbf{I}) \\ & \frac{\partial \mathcal{F}}{\partial \boldsymbol{\lambda}} = \mathbb{E}_{q_0} \nabla_{\boldsymbol{\lambda}} \log p(\mathbf{z}|\mathbf{z}) - \mathbb{E}_{q_0} \nabla_{\boldsymbol{\lambda}} \log q(\mathbf{z}|\boldsymbol{\lambda}) + \mathbb{E}_{q_0} \nabla_{\boldsymbol{\lambda}} \log \hat{p}(\mathbf{z}|\boldsymbol{\tau}) \\ & \frac{\partial \mathcal{F}}{\partial \boldsymbol{\tau}} = \mathbb{E}_{q_0} \nabla_{\boldsymbol{\tau}} \log \hat{p}(\mathbf{z}|\boldsymbol{\tau}) - \mathbb{E}_{q_0} \nabla_{\boldsymbol{\tau}} \log Z(\boldsymbol{\tau}) \end{split} \tag{2}$$

- ▶ $\log p(x|z)$ \log -likelihood of our model
- ▶ $\log \hat{p}(\mathbf{z}|\mathbf{\tau}) = \sum_{i} \tau_{i} R_{i}(\mathbf{z})$ regularization term.

- $\blacktriangleright \ \mathbb{E}_{q_0} \nabla_{\boldsymbol{\tau}} \log \hat{p}(\mathbf{z}|\boldsymbol{\tau}) = \mathbb{E}_{q_0} \mathbf{R}(\mathbf{z})$
- lacksquare $abla_{ au} \log Z(oldsymbol{ au}) = \mathbb{E}_{oldsymbol{
 ho}(oldsymbol{z}|oldsymbol{ au})} oldsymbol{R}(oldsymbol{z})$

How to optimize with $Z(\tau)$ unknown?

$$\nabla_{\tau} \log Z(au) = \mathbb{E}_{p(\mathbf{z}| au)} R(\mathbf{z})$$

We need samples from $p(\mathbf{z}|\boldsymbol{\tau})$

- ► Contrastive divergence
 - After sampling $z_i \sim q_\lambda$ do **a few** iterations of MCMC (starting from z_i) to obtain samples from $p(z|\tau)$.
- Importance sampling
 - $Z(\tau) = \mathbb{E}_{q_0} \frac{\hat{p}(z|\tau)}{q_0(z)}$
 - $\blacktriangleright \mathbb{E}_{p(\boldsymbol{z}|\boldsymbol{\tau})}R(\boldsymbol{z}) = \frac{1}{Z(\boldsymbol{\tau})}\mathbb{E}_{q_0}\frac{R(\boldsymbol{z})\hat{p}(\boldsymbol{z}|\boldsymbol{\tau})}{q_0(\boldsymbol{z})}$
 - We use the same samples $z_i \sim q_\lambda$ to estimate Z(au) and $\mathbb{E}_{q_0} \frac{R(z)\hat{p}(z| au)}{q_0(z)}$

Review

Our algorithm

- Our algorithm is scalable on data size
- ▶ Time of each iteration on one batch of D docs with N total words is O(N + TD + TW) (no DW term)
- ...or $O(N + T^2D + TW)$ with decorrelation
- ▶ Returns both fitted regularization coeffitients τ and matricies (Φ,Θ) whole distribution over (Φ,Θ) .
- ▶ Coefficients τ and distributions (Φ, Θ) are fitted **simultaneously** as long as algorithms iterates over data (over batches)

Questions still not answered

- What to do with that distribuition (what is the final answer)
 - Sample from it (probably multiple times)
 - Return median $\mathcal{P}(\mu)$
 - Compute mode $\operatorname{argmax}_{\Phi,\Theta} q(\Phi,\Theta|\boldsymbol{\lambda})$
- Is maximum-evidence approach really going to work in this scenario?
- How to organize online-learning process carefully?

Thank you for your attention!